

When size matters: attention affects performance by contrast or response gain

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Covert attention, the selective processing of visual information in the absence of eye movements, improves behavioral performance. We found that attention, both exogenous (involuntary) and endogenous (voluntary), can affect performance by contrast or response gain changes, depending on the stimulus size and the relative size of the attention field. These two variables were manipulated in a cueing task while stimulus contrast was varied. We observed a change in behavioral performance consonant with a change in contrast gain for small stimuli paired with spatial uncertainty and a change in response gain for large stimuli presented at one location (no uncertainty) and surrounded by irrelevant flanking distracters. A complementary neuroimaging experiment revealed that observers' attention fields were wider with than without spatial uncertainty. Our results support important predictions of the normalization model of attention and reconcile previous, seemingly contradictory findings on the effects of visual attention.

There has been a long-standing debate about the neural computations underlying selective visual attention^{1–5}. Experiments examining how covert attention modulates visual performance and neuronal activity in visual cortex have documented a variety of empirical phenomena, some of which appear to be mutually contradictory. Central to this debate are studies of the interactions between attention and stimulus contrast. Some results suggest that attention increases neuronal responses by a multiplicative response gain factor⁶ (**Fig. 1a**) and similarly improves performance via upward scaling of the psychometric function^{7–12}. Other neurophysiological and behavioral results suggest a change in contrast gain (**Fig. 1b**), that is, a leftward shift of the contrast-response^{2,13–15} and psychometric^{9,10} functions. Still other results suggest an additive effect of attention across the entire contrast range or a combination of both response gain and contrast gain changes^{16–18}. These ostensibly contradictory findings have been taken to represent alternative models of attention^{2,4,6,10,14,15,19}.

The normalization model of attention¹ was proposed to reconcile these (and other) seemingly conflicting, empirical findings of the effects of attention on sensory responses in visual cortex. This computational theory identifies two critical factors that determine the effect of attention on contrast-response functions and, consequently, on behavioral psychometric functions: the size of the stimulus and the spatial spread of attention, known as the attention field. Changing the relative sizes of these two factors allows the model to exhibit response gain changes, contrast gain changes, and various combinations of response gain and contrast gain changes. Specifically, the model predicts that attention increases response gain when the stimulus is large and the attention field small, and increases contrast gain when the stimulus is small and the attention field is large (**Fig. 1**)¹. The core idea is that the attention field reshapes the distribution of activity across the population of neurons, shifting the balance between excitation and suppression, yielding either a change in response gain, contrast

gain or a combination of the two (see ref. 1 for mathematical derivation). Note that changing the size and shape of the attention field on its own would not predict a shift between response gain changes and contrast gain changes. Instead, it is the architecture of the model that makes this prediction, specifically because the effect of attention is to multiply the stimulus-evoked activity before divisive normalization¹. Alternatively, if attention modulated activity after normalization then it would always yield response gain changes, regardless of the size or shape of the attention field.

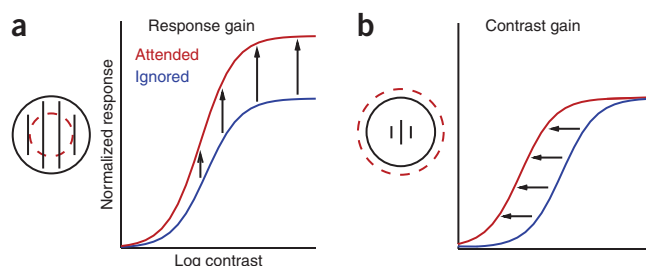
Other computational models of attention, although apparently similar to the normalization model of attention¹, do not predict the shift from response to contrast gain changes. Some of these models presume that spatial attention always has the same effect on the excitation and suppression (that is, the numerator and the denominator of the normalization equation), always yielding a contrast gain change^{2,4,20}. In another model, attention affects only the strength of the normalization, always yielding a response gain change¹⁹. Unique to the normalization model of attention is the idea that the effects of attention on the numerator and denominator can differ depending on the relative sizes of the stimulus and the attention field, altering the balance between excitation and suppression.

We empirically tested the normalization model of attention¹. Specifically, we tested the key prediction that the effect of attention can systematically shift from a change in response gain to contrast gain with smaller stimuli and a broader attention field. There is evidence that the attention field is measurable, flexible in size (within limits) and can be experimentally manipulated^{21–24}. Previous studies, however, have not systematically manipulated the size of the attention field and its relation to stimulus size. We used spatial uncertainty to manipulate attention field size and confirmed, in a complementary experiment using functional magnetic resonance imaging (fMRI), that the attention field was larger with spatial uncertainty than without it.

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Figure 1 Normalization model of attention exhibits qualitatively different forms of attentional modulation, depending on stimulus size and attention field size (adapted from ref. 1). Each panel shows contrast-response functions for a simulated neuron when attending to a stimulus in the neuron's receptive field and when attending to a stimulus in the opposite hemifield. **(a)** Response gain (largest effects at higher contrasts, upward shift of the contrast-response function) for large stimulus size and small attention field. **(b)** Contrast gain (largest effect at intermediate contrasts, appears as a leftward shift of the contrast-response function) for small stimulus size and large attention field. Red, simulated responses as a function of contrast when stimuli in the receptive field were attended. Blue, simulated responses when attending to the opposite hemifield. Only stimulus size and attention field size were changed in simulations; all other model parameters were identical in both panels. The solid black circle indicates simulated receptive field size. The dashed red circle indicates simulated attention field size. The vertical black grating indicates stimulus size.



We adopted established experimental protocols to investigate two types of attention: exogenous and endogenous^{7,25–29}. Exogenous attention is involuntary, stimulus driven and has a transient effect, which peaks at about 100 ms and decays shortly thereafter. Endogenous attention is voluntary, conceptually driven (for example, according to instructions) and has a sustained effect, which takes about 300 ms to be deployed and can last up to seconds. These two types of attention have been shown to have distinct effects on behavioral performance^{9,10,26,27,29,30}. We found, for both exogenous and endogenous attention, that the effects of attention on behavioral performance could be shifted from one type of gain modulation to the other by manipulating just the stimulus size and attention field size.

RESULTS

Psychophysics: response gain or contrast gain

Experimental parameters were optimized (Online Methods) to maximize the effects of the two types of attention: exogenous (**Fig. 2a**) and endogenous (**Fig. 2b**)^{7–9,27,29,31}. In both experiments, observers performed an orientation discrimination task on one of two grating patches; for each of several contrasts (the contrasts of both gratings were identical on any given trial and covaried across trials in random order). Covert attention (without eye movements) was directed by a pre-cue. A response-cue at stimulus offset indicated the target location, yielding three cue conditions: valid (pre-cue matched response-cue), invalid (mismatched) and neutral. Comparing performance accuracy (d') for valid, neutral and invalid trials revealed differences in behavioral performance: benefits at the cued location and costs at the uncued location compared with the neutral condition. To enlarge the attention field, we presented stimuli with spatial uncertainty (one of five locations randomly interleaved; **Fig. 2c**). To narrow the attention

field, we presented stimuli at a fixed location; for endogenous attention, they were also surrounded with irrelevant, flanking distracters. Smaller stimuli were paired with spatial uncertainty (**Fig. 2c**) and larger stimuli were paired with no uncertainty (**Fig. 2d**).

Both exogenous and endogenous attention exerted different gain changes on psychometric functions (either contrast gain or response gain changes), depending on stimulus size and uncertainty. The measured psychometric functions were fit (nonlinear least squares, Online Methods) with a parametric function. The two parameters d'_{\max} (asymptotic performance at high contrast levels) and c_{50} (the contrast yielding half maximum performance) determined response gain and contrast gain, respectively. They were estimated separately for each attention condition, whereas the exponent n (slope) was constrained to have the same value for all conditions.

When the targets were large and the spatial uncertainty was minimal (encouraging deployment of a relatively small attention field), attention increased asymptotic performance at high contrasts, as predicted by a response gain change (robust differences in d'_{\max} , but no evidence for a change in c_{50} ; **Fig. 3a,b** and **Supplementary Tables 1** and **2**). Conversely, when the targets were small and the attention field was relatively large as a result of spatial uncertainty, attention altered performance primarily for intermediate contrasts, as predicted by a contrast gain change (reliable differences in c_{50} ; **Fig. 3c,d** and **Supplementary Tables 1** and **2**). There was a trend for a change in d'_{\max} that was not statistically significant (see **Supplementary Tables 1** and **2**); this trend was not inconsistent with the predictions of the normalization model of attention, as the model predicts a pure contrast gain change only in the limit as the attention field size increases. Best-fitting values for the exponents of the psychometric functions were similar for the exogenous ($n = 2.48$, 68% confidence interval = [2.37, 2.66]) and endogenous ($n = 2.00$, 68% confidence interval = [1.85, 2.22]) attention experiments (**Supplementary Fig. 1**), and were in the range of values reported for contrast-response functions in early visual cortex³². We refit the data, allowing the exponent to vary independently across conditions, and observed the same pattern of results: a response gain change for large stimuli with no uncertainty and a contrast gain change for small stimuli with spatial uncertainty.

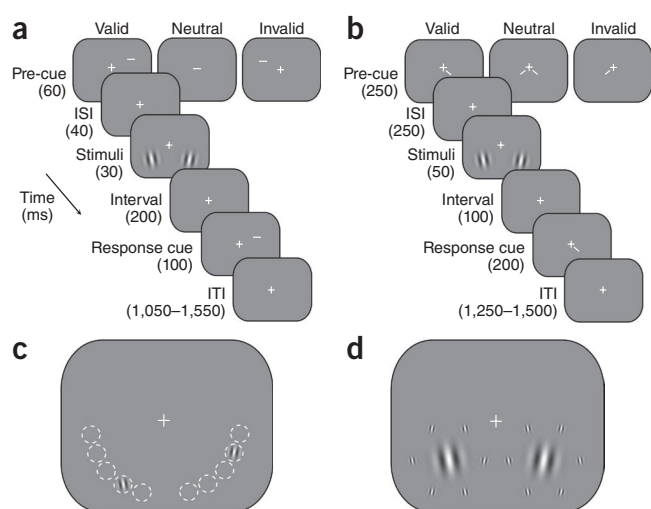
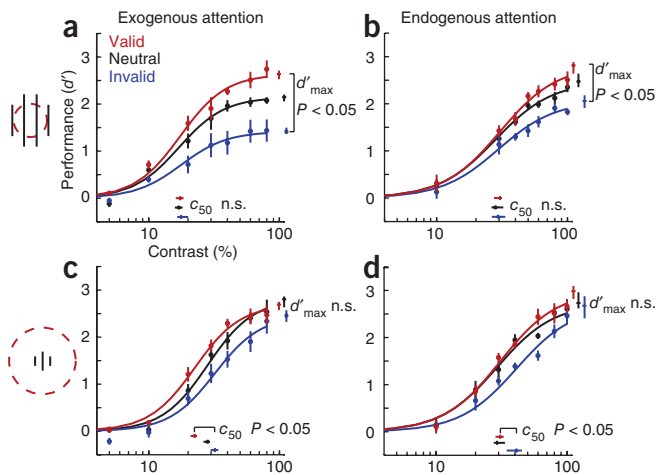


Figure 2 Experimental protocols. **(a)** Exogenous attention task. ISI, interstimulus interval; ITI, intertrial interval. **(b)** Endogenous attention task. **(c)** Small stimuli were presented with spatial uncertainty, at one of five predefined isoecentric locations. Across trials, stimulus locations varied randomly and independently on the left and right sides to encourage observers to employ a larger attention field. The dashed, white circles indicate the possible stimulus locations (not displayed during the experiments). **(d)** Large stimuli were presented at fixed stimulus locations with no spatial uncertainty (centered at the middle of the five locations of **c**). To narrow the size of the attention field for endogenous attention, the two large Gabor stimuli were each surrounded by six irrelevant distracters.



The same results were evident in the psychometric functions from individual observers (Fig. 4). Large stimuli with small attention fields yielded similar c_{50} values for valid and invalid pre-cues (one-tailed Wilcoxon, $P = 0.500$; Fig. 4a), whereas small stimuli with large attention fields consistently resulted in smaller c_{50} values for valid pre-cues (one-tailed Wilcoxon, $P = 0.007$; Fig. 4a). Large stimuli with small attention fields reliably resulted in larger d'_{\max} values for valid than for invalid pre-cues (one-tailed Wilcoxon, $P < 0.001$; Fig. 4b), whereas small stimuli with large attention fields resulted in similar d'_{\max} values for the two pre-cue conditions (one-tailed Wilcoxon, $P = 0.221$; Fig. 4b).

As a complementary means of testing the predictions of the model, we used nested hypothesis tests (nested F tests) to compare the full model fit with four restricted models: normalization model prediction (constraining c_{50} to be constant across cue conditions (valid, neutral, invalid) for large stimuli without uncertainty and constraining d'_{\max} for small stimuli with uncertainty), opposite of the normalization model prediction (constraining d'_{\max} for large stimuli without uncertainty and constraining c_{50} for small stimuli with uncertainty), always response gain (constraining c_{50} for both experimental conditions) and always contrast gain (constraining d'_{\max} for both experimental conditions). As predicted by the normalization model of attention, the full model fit the mean data (across observers) similar to the first model (exogenous $P > 0.1$ and endogenous $P > 0.1$) and fit better than any of the other restricted models (second model, exogenous $P < 0.001$ and endogenous $P = 0.002$; third model, exogenous $P = 0.035$ and endogenous $P > 0.1$; fourth model, exogenous $P < 0.001$ and endogenous $P = 0.006$). Only one of these eight nested hypothesis tests failed to reject an alternative to the normalization model; for endogenous attention, the full model fit better than the third model, as predicted by the normalization model, but it was not statistically significant ($P > 0.1$). A nonparametric cross-validation analysis of the model fits yielded results consistent with the nested F tests. In short, for exogenous attention, we were able to reject contrast gain

Figure 4 Effects of stimulus and attention field size on parameter estimates of individual observers. (a) Contrast yielding half maximum performance (c_{50}), estimated for valid pre-cues versus invalid pre-cues. (b) Asymptotic performance at high contrast (d'_{\max}). Red symbols indicate large stimulus with small attention field. Blue symbols indicate small stimulus with large attention field. Open symbols indicate individual observers. Filled symbols indicate mean across observers. Squares indicate exogenous attention. Circles indicate endogenous attention. Observer 1 participated in both the exogenous and endogenous attention experiments and is marked with diagonal lines through the symbols.

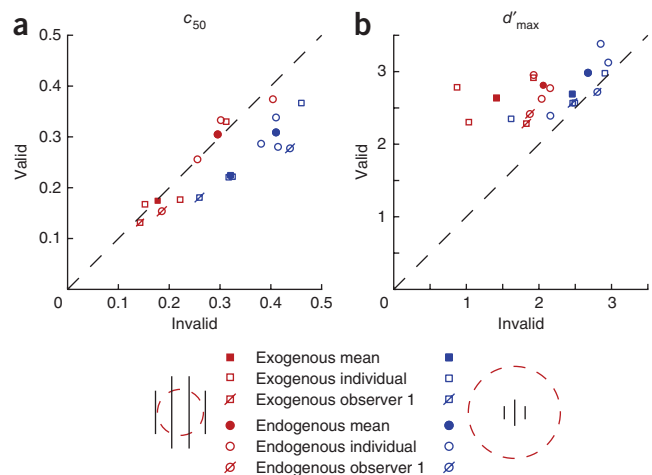
Figure 3 Effects of exogenous and endogenous attention on performance (d') as a function of contrast. (a,b) Large stimulus with small attention field. (c,d) Small stimulus with large attention field. Exogenous attention is shown in a and c. Endogenous attention is shown in b and d. Shown are plots of psychometric functions for each attentional condition (valid, neutral and invalid pre-cues) and parameter estimates (c_{50} , contrast yielding half-maximum performance; d'_{\max} , asymptotic performance at high contrast). Exponent n (slope) was constrained to have the same value for all pre-cue conditions. Each data point represents the mean across observers. Error bars on data points are \pm s.e.m. ($n = 4$ observers in each experiment, 9,408–14,910 trials per observer, except one observer who completed 21,000 in the exogenous experiment). Error bars on parameter estimates are 68% confidence intervals, obtained by bootstrapping.

only and response gain only in favor of the normalization model; for endogenous attention, we were able to reject contrast gain only, but did not have enough statistical power to reject response gain only. As stated above, this is not inconsistent with the predictions of the normalization model of attention, as the model predicts a pure contrast gain change only in the limit as the attention field size increases.

Fixation was stable during stimulus presentation in the exogenous and endogenous attention experiments (Supplementary Fig. 2). The recorded gaze positions of all observers had a standard deviation of 0.185° horizontally and 0.300° vertically in the exogenous experiment, and 0.257° horizontally and 0.426° vertically in the endogenous attention experiment. Gaze position distributions (horizontal and vertical) were statistically indistinguishable for all conditions (leftward, rightward and neutral pre-cues, with or without spatial uncertainty; all two-tailed t test comparisons $P > 0.01$; two comparisons at $P > 0.06$; Bonferroni critical level = 0.004) in both experiments. Saccades were detected in 0.113% and 0.749% of the trials in the exogenous and endogenous attention experiments, respectively.

fMRI: attention field size

The size of the attention field was larger when stimuli were presented with spatial uncertainty than when the stimuli were presented at the same location without spatial uncertainty (Fig. 5). We used fMRI to measure attention field size while observers performed the endogenous attention task in the MRI scanner (Online Methods). We analyzed activity in primary visual cortex (V1), as opposed to other visual areas (cortical or subcortical), as V1 is known to process contrast and orientation, attention robustly modulates fMRI responses in V1 (refs. 28,33–37) and concurrently improves behavior^{28,34}, and V1 is relatively large compared with other visual areas, with a precise retinotopic map, enabling us to more easily quantify the spatial extent



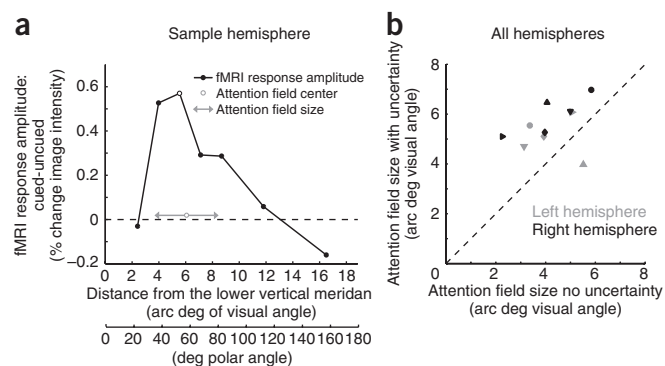


Figure 5 Attention field size depends on spatial uncertainty. **(a)** fMRI responses at a series of locations in left V1 for a typical observer when a small stimulus was presented at the center of the five possible locations during the experiment with spatial uncertainty. The upper abscissa indicates the distance along a 6° isoeccentric arc in the visual field that included the stimulus location. Zero corresponds to the location at which the arc crosses the lower vertical meridian. The lower abscissa indicates the distance in degrees of polar angle. Black symbols represent fMRI response differences (cued minus uncued), plotted at locations corresponding to the centroids of each of several V1 regions of interest, defined based on retinotopic mapping. Open symbol, location corresponding to the center of the stimulus. The gray symbol represents the estimated center of the attention field. The gray line represents the estimated spread of the attention field. **(b)** fMRI measurements of the spread of cortical activity in V1, in units of arc degrees of visual angle in the visual field, with spatial uncertainty versus without spatial uncertainty. Each symbol shape corresponds to one observer. Gray indicates left hemisphere. Black indicates right hemisphere.

of activation. To evaluate whether spatial uncertainty encouraged observers to enlarge their attention field, we compared the spread of cortical activity brought about by attention in two conditions: spatial uncertainty and no spatial uncertainty. In one condition, a pair of stimuli (10% contrast) appeared in the left and right hemifields, randomly and independently at one of five predefined locations in each hemifield. In the second condition, stimuli of the same size and contrast were always shown at a fixed pair of locations (the middle of the five locations in the first condition) in the left and right hemifields, with placeholders continuously indicating these two locations.

fMRI response amplitudes were measured at each of several locations in V1 corresponding retinotopically to an arc in the visual field that included the five stimulus locations (Online Methods). Data were analyzed from the first (spatial uncertainty) condition only for trials on which the stimuli were presented at the middle of the five locations. Thus, responses were measured for stimuli of identical size at the identical stimulus location, but with and without spatial uncertainty. We then subtracted the response amplitudes for uncued stimulus presentations from the response amplitudes for cued stimuli presented at the same stimulus location, largely removing any stimulus-evoked response. This resulted in a spatial distribution of response differences; that is, the attention field for each condition (with and without spatial uncertainty), separately for left and right hemisphere V1 from each observer. Finally, attention field size was quantified as the spread of this spatial distribution of response differences (Fig. 5a) and was found to be larger when stimuli were presented with spatial uncertainty than without it (one-tailed Wilcoxon test, $P = 0.006$; Fig. 5b).

Eye position was measured during the fMRI experiment and, as in the behavioral experiments, fixation was stable (Supplementary Fig. 3). We analyzed eye movements for the spatial uncertainty condition only for trials in which stimuli were presented at the middle of the five

locations. Recorded gaze positions of all observers during stimulus presentation had a standard deviation of 0.326° horizontally and 0.443° vertically with spatial uncertainty, and 0.438° horizontally and 0.532° vertically without spatial uncertainty. Gaze position distributions (horizontal and vertical) were statistically indistinguishable for all conditions (leftward and rightward pre-cues, with or without spatial uncertainty; all two-tailed t test comparisons $P > 0.1$) and no saccades were detected.

DISCUSSION

The interpretation of the behavioral data is based on a model of the noise or variability in neuronal activity that limits behavioral performance along with a model of the decision-making process^{10,38}. Performance accuracy, d' , is proportional to the signal-to-noise ratio of the underlying neuronal responses. Thus, with additive, independent, and identically distributed noise and a maximum-likelihood decision rule, any change in the neuronal contrast-response functions would be reflected as a parallel change in performance accuracy. A change in response gain of the underlying neuronal responses (Fig. 1a) would yield a scaling of the psychometric function (Fig. 3a,b) and a change in contrast gain of the underlying neuronal responses (Fig. 1b) would yield a horizontal shift (on the log contrast axis) of the psychometric function (Fig. 3c,d). This interpretation of our behavioral results depends on the assumption that behavioral performance is limited by additive independent, and identically distributed noise, and although there are psychophysical data suggesting that this may be the case³⁹, we must also consider the possibility that performance is instead limited by the Poisson-like noise evident in single-cell firing rates^{40,41}. But this alternative model of the noise, coupled with a maximum-likelihood decision rule³⁸, leads to exactly the same interpretation of the psychometric functions¹⁰.

The exogenous attention task may have involved some degree of endogenous attention as well. The advantage of using an exogenous cue is that it results in a highly controlled manipulation of attention that is not dependent on the observer's strategy^{7,8,27,29}. However, observers knew before each trial that the discrimination task was to be performed on one of two stimuli and could have therefore directed endogenous attention to both, and it is not known exactly how exogenous attention (driven by the pre-cue) and endogenous attention (driven by the observer's knowledge of the task) combine. One possibility, which is theoretically tractable and consistent with our results, is that they combine multiplicatively. An alternative possibility is that exogenous attention completely overrides endogenous attention; our results suggest that this is not the case, as adding spatial uncertainty shifted performance from a response gain change to a contrast gain change.

Spatial uncertainty was used to manipulate attention field size, but we ensured that this manipulation did not affect task difficulty and performance accuracy. Specifically, we adjusted the degree of stimulus tilt separately for each observer and for each of the two stimulus and attention field sizes so that performance was approximately 85% correct at full contrast with a neutral cue. Moreover, we did not compare performance between conditions with different amounts of uncertainty. Instead, within each condition (small stimuli with large uncertainty, large stimuli with no uncertainty), we compared performance for valid, invalid and neutral cues; thus, there was no difference in spatial uncertainty in each of these comparisons.

We used fMRI to measure attention field size by taking advantage of the fact that neuromodulatory inputs contribute to the fMRI responses. The attention field, according to the normalization model of attention, is hypothesized to be a neuromodulatory input, largely invisible in extracellular electrophysiological measurements of

spiking activity. Although it is known that the fMRI signal is triggered by metabolic demands of neural activity, the details of this process are only partially understood, and it is likely that a mixture of synaptic and spiking activity may contribute^{42,43}. The extent of decoupling between synaptic and spiking activity depends on the nature of cortical processing, that is, whether the cortical activity is dominated by local recurrent circuitry or by synaptic inputs to a cortical area (either feedforward or feedback) from other brain areas. Consequently, fMRI measurements may be highly correlated with the output spiking activity of a visual area under some circumstances, but with subthreshold neuromodulatory input under other circumstances. Indeed, strong neuromodulatory input as a result of attention might explain the apparent discrepancy in the literature between attentional effects measured with fMRI and electrophysiologically. The implication is that fMRI is not merely a low-resolution, noninvasive measure of spiking activity, but instead it constitutes an independent, complementary measure⁴⁴.

We measured attention field size using low-contrast stimuli with high attentional demands to put the cortical circuits into a regime in which cortical activity is dominated by feedback synaptic inputs. The spread of cortical activity was substantially larger with spatial uncertainty than with no uncertainty. These results provide converging evidence for the idea that manipulating spatial uncertainty can affect attention field size^{21–24,45,46}. Future studies should use behavioral and neuroimaging data, acquired simultaneously, to further test predictions of the normalization model of attention¹; for example, by simultaneously fitting measurements of behavioral performance and attention field size or by binning trials according to attention field size and testing for systematic differences in behavioral performance.

Previous studies have reported ostensibly conflicting results with regard to the gain changes induced by selective attention. A review of the literature¹ suggests that discrepancies in the neurophysiological findings may have resulted from differences in the experimental protocols, specifically the relative sizes of the stimuli and attention fields, which have not been systematically controlled and manipulated. On the basis of our results, we suggest that such differences in the experimental protocols may also explain previous discrepancies among psychophysical studies^{7–12,18}. For instance, with constant stimulus size, some studies have found that exogenous attention alters performance via a response gain change, whereas endogenous attention does so via a contrast gain change^{7–10}. A response gain change could have been elicited by brief peripheral cues nearby the stimulus, whereas a contrast gain change could have resulted from endogenous cues at fixation, rather than cues adjacent to the stimulus, which may have encouraged a narrower or larger attention field, respectively. We were able to manipulate the stimulus size and the size of the attention field with both types of attention to yield either contrast gain or response gain changes.

We conclude that attention modulates activity in visual cortex in a manner that can resemble either a change in response gain or contrast gain, depending on stimulus size and attention field size. Our results provide, to the best of our knowledge, the first experimental evidence supporting important predictions of the normalization model of attention, thereby furthering our understanding of the processing in visual cortex and the neural computations underlying visual attention.

METHODS

Methods and any associated references are available in the online version of the paper at <http://www.nature.com/natureneuroscience/>.

Note: Supplementary information is available on the Nature Neuroscience website.

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AUTHOR CONTRIBUTIONS

K.H. programmed, conducted and analyzed the experiments and co-wrote the manuscript. L.M.-K. conducted and analyzed the psychophysics experiments and assisted in conducting and programming the fMRI experiment. M.C. and D.J.H. conceived and supervised the project and co-wrote the manuscript.

COMPETING FINANCIAL INTERESTS

The authors declare no competing financial interests.

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1. Reynolds, J.H. & Heeger, D.J. The normalization model of attention. *Neuron* **61**, 168–185 (2009).
2. Reynolds, J.H. & Chelazzi, L. Attentional modulation of visual processing. *Annu. Rev. Neurosci.* **27**, 611–647 (2004).
3. Carrasco, M. Covert attention increases contrast sensitivity: psychophysical, neurophysiological and neuroimaging studies. *Prog. Brain Res.* **154**, 33–70 (2006).
4. Boynton, G.M. A framework for describing the effects of attention on visual responses. *Vision Res.* **49**, 1129–1143 (2009).
5. Desimone, R. & Duncan, J. Neural mechanisms of selective visual attention. *Annu. Rev. Neurosci.* **18**, 193–222 (1995).
6. McAdams, C.J. & Maunsell, J.H. Effects of attention on orientation-tuning functions of single neurons in macaque cortical area V4. *J. Neurosci.* **19**, 431–441 (1999).
7. Pestilli, F. & Carrasco, M. Attention enhances contrast sensitivity at cued and impairs it at uncued locations. *Vision Res.* **45**, 1867–1875 (2005).
8. Pestilli, F., Viera, G. & Carrasco, M. How do attention and adaptation affect contrast sensitivity? *J. Vis.* **7**, 9 1–12 (2007).
9. Ling, S. & Carrasco, M. Sustained and transient covert attention enhance the signal via different contrast response functions. *Vision Res.* **46**, 1210–1220 (2006).
10. Pestilli, F., Ling, S. & Carrasco, M. A population-coding model of attention's influence on contrast response: estimating neural effects from psychophysical data. *Vision Res.* **49**, 1144–1153 (2009).
11. Morrone, M.C., Denti, V. & Spinelli, D. Color and luminance contrasts attract independent attention. *Curr. Biol.* **12**, 1134–1137 (2002).
12. Morrone, M.C., Denti, V. & Spinelli, D. Different attentional resources modulate the gain mechanisms for color and luminance contrast. *Vision Res.* **44**, 1389–1401 (2004).
13. Li, X., Lu, Z.L., Tjan, B.S., Doshier, B.A. & Chu, W. Blood oxygenation level-dependent contrast response functions identify mechanisms of covert attention in early visual areas. *Proc. Natl. Acad. Sci. USA* **105**, 6202–6207 (2008).
14. Martínez-Trujillo, J. & Treue, S. Attentional modulation strength in cortical area MT depends on stimulus contrast. *Neuron* **35**, 365–370 (2002).
15. Reynolds, J.H., Pasternak, T. & Desimone, R. Attention increases sensitivity of V4 neurons. *Neuron* **26**, 703–714 (2000).
16. Buracas, G.T. & Boynton, G.M. The effect of spatial attention on contrast response functions in human visual cortex. *J. Neurosci.* **27**, 93–97 (2007).
17. Williford, T. & Maunsell, J.H. Effects of spatial attention on contrast response functions in macaque area V4. *J. Neurophysiol.* **96**, 40–54 (2006).
18. Huang, L. & Dobkins, K.R. Attentional effects on contrast discrimination in humans: evidence for both contrast gain and response gain. *Vision Res.* **45**, 1201–1212 (2005).
19. Lee, J. & Maunsell, J.H. A normalization model of attentional modulation of single unit responses. *PLoS ONE* **4**, e4651 (2009).
20. Reynolds, J.H., Chelazzi, L. & Desimone, R. Competitive mechanisms subserve attention in macaque areas V2 and V4. *J. Neurosci.* **19**, 1736–1753 (1999).
21. Datta, R. & DeYoe, E.A. I know where you are secretly attending! The topography of human visual attention revealed with fMRI. *Vision Res.* **49**, 1037–1044 (2009).
22. Eriksen, C.W. & St James, J.D. Visual attention within and around the field of focal attention: a zoom lens model. *Percept. Psychophys.* **40**, 225–240 (1986).
23. Müller, N.G., Bartelt, O.A., Donner, T.H., Villringer, A. & Brandt, S.A. A physiological correlate of the “Zoom Lens” of visual attention. *J. Neurosci.* **23**, 3561–3565 (2003).
24. Castiello, U. & Umiltà, C. Size of the attentional focus and efficiency of processing. *Acta Psychol. (Amst.)* **73**, 195–209 (1990).
25. Ling, S. & Carrasco, M. When sustained attention impairs perception. *Nat. Neurosci.* **9**, 1243–1245 (2006).
26. Lu, Z.L. & Doshier, B.A. Spatial attention: different mechanisms for central and peripheral temporal precues? *J. Exp. Psychol. Hum. Percept. Perform.* **26**, 1534–1548 (2000).

27. Nakayama, K. & Mackeben, M. Sustained and transient components of focal visual attention. *Vision Res.* **29**, 1631–1647 (1989).
28. Liu, T., Pestilli, F. & Carrasco, M. Transient attention enhances perceptual performance and fMRI response in human visual cortex. *Neuron* **45**, 469–477 (2005).
29. Giordano, A.M., McElree, B. & Carrasco, M. On the automaticity and flexibility of covert attention: a speed-accuracy trade-off analysis. *J. Vis.* **9**, 30 31–10 (2009).
30. Yeshurun, Y., Montagna, B. & Carrasco, M. On the flexibility of sustained attention and its effects on a texture segmentation task. *Vision Res.* **48**, 80–95 (2008).
31. Jonides, J. & Irwin, D.E. Capturing attention. *Cognition* **10**, 145–150 (1981).
32. Sclar, G., Maunsell, J.H. & Lennie, P. Coding of image contrast in central visual pathways of the macaque monkey. *Vision Res.* **30**, 1–10 (1990).
33. Kastner, S., Pinsk, M.A., De Weerd, P., Desimone, R. & Ungerleider, L.G. Increased activity in human visual cortex during directed attention in the absence of visual stimulation. *Neuron* **22**, 751–761 (1999).
34. Gandhi, S.P., Heeger, D.J. & Boynton, G.M. Spatial attention affects brain activity in human primary visual cortex. *Proc. Natl. Acad. Sci. USA* **96**, 3314–3319 (1999).
35. Tootell, R.B. *et al.* The retinotopy of visual spatial attention. *Neuron* **21**, 1409–1422 (1998).
36. Brefczynski, J.A. & DeYoe, E.A. A physiological correlate of the 'spotlight' of visual attention. *Nat. Neurosci.* **2**, 370–374 (1999).
37. Somers, D.C., Dale, A.M., Seiffert, A.E. & Tootell, R.B. Functional MRI reveals spatially specific attentional modulation in human primary visual cortex. *Proc. Natl. Acad. Sci. USA* **96**, 1663–1668 (1999).
38. Jazayeri, M. & Movshon, J.A. Optimal representation of sensory information by neural populations. *Nat. Neurosci.* **9**, 690–696 (2006).
39. Katkov, M., Tsodyks, M. & Sagi, D. Inverse modeling of human contrast response. *Vision Res.* **47**, 2855–2867 (2007).
40. Dean, A.F. The variability of discharge of simple cells in the cat striate cortex. *Exp. Brain Res.* **44**, 437–440 (1981).
41. Carandini, M. Amplification of trial-to-trial response variability by neurons in visual cortex. *PLoS Biol.* **2**, e264 (2004).
42. Heeger, D.J. & Ress, D. What does fMRI tell us about neuronal activity? *Nat. Rev. Neurosci.* **3**, 142–151 (2002).
43. Logothetis, N.K. & Wandell, B.A. Interpreting the BOLD signal. *Annu. Rev. Physiol.* **66**, 735–769 (2004).
44. Bartels, A., Logothetis, N.K. & Moutoussis, K. fMRI and its interpretations: an illustration on directional selectivity in area V5/MT. *Trends Neurosci.* **31**, 444–453 (2008).
45. Hernandez, M., Costa, A. & Humphreys, G.W. The size of an attentional window affects working memory guidance. *Atten. Percept. Psychophys.* **72**, 963–972.
46. Belopolsky, A.V., Zwaan, L., Theeuwes, J. & Kramer, A.F. The size of an attentional window modulates attentional capture by color singletons. *Psychon. Bull. Rev.* **14**, 934–938 (2007).



ONLINE METHODS

Experiments were conducted with the written consent of each observer and procedures were approved by the University Committee on Activities Involving Human Subjects at New York University.

Observers and psychophysical sessions. Seven observers (25–40 years old, five females) with normal or corrected vision participated in the experiments. Half were assigned to the exogenous and the other half to the endogenous attention experiment (one participated in both). For each experiment, observers completed both conditions (small stimuli paired with large attention field and vice versa) in a counter-balanced order. For each condition, observers participated in practice sessions to determine individual orientation discrimination thresholds, followed by four to six 1-h experimental sessions.

Exogenous attention task. On each trial, a pair of Gabor stimuli (5 cycles per degree) were presented in the lower quadrants of the left and right hemifields at 5° eccentricity, one of which was the target. Contrast varied from trial to trial in randomly shuffled order and stimuli were presented briefly to avoid any possible dependence of attentional state on stimulus contrast. Observers performed an orientation discrimination task to indicate whether the 30-ms target was tilted left or right of vertical and received auditory feedback if their response was incorrect. On two-thirds of the trials, exogenous attention was directed by flashing the 60-ms pre-cue (0.5° white line) above one of the stimulus locations, but not at the stimulus location to avoid masking. On the remaining trials, a neutral pre-cue was presented at fixation, providing a baseline against which to evaluate the effects of attention. A 40-ms interstimulus interval between cue offset and stimulus onset, yielding a cue-stimulus onset asynchrony of 100 ms, maximized the effect of exogenous attention^{27,29,31}. Target location was indicated by a peripheral 100-ms response-cue (0.5° green line). A valid cue was defined as a match between the pre-cue location and the response-cue location (one-third of the trials); a mismatch yielded an invalid cue (one-third of the trials). Observers were explicitly told that pre-cues were randomized and uninformative about the target location.

Endogenous attention task. The endogenous attention task was carried out similar to the exogenous attention task, except for the following modifications. Durations of the central pre-cues (250 ms) and interstimulus intervals (250 ms) were increased to ensure that observers had enough time to deploy endogenous attention^{27,29,47}. Stimuli were presented at 6° eccentricity. A 200-ms response-cue appeared right or left of the fixation cross. One-fourth of the trials had neutral cues. Of the remaining trials, two-thirds were valid and one-third were invalid to ensure that observers followed the cue^{25,29}. Observers were explicitly told that the cue was informative regarding the target location and that there was a benefit in using the cue to perform the task. Using a partially valid cue allowed us to assess both the benefit and cost of attention at the attended and unattended locations, respectively, even though observers are known to treat the ~70% valid cue much as they would have treated a cue with 100% validity⁴⁸.

Psychophysical manipulations of the stimulus and attention field size. Stimulus size and attention field size covaried systematically in separate sessions. Small stimuli ($\sigma = 0.4^\circ$, s.d. of Gabor window) were presented at one of five predefined isoeccentric locations in each of the lower right and left quadrants. Across trials, stimulus locations varied randomly and independently on the left and right sides. Peripheral pre-cue locations covaried with the stimulus locations in the exogenous attention experiment. Large stimuli ($\sigma = 1^\circ$) were presented at fixed locations (same eccentricity as small stimuli) with no spatial uncertainty. In the endogenous attention experiment, the two large Gabor stimuli were each surrounded by six irrelevant distracters (Gabor flankers, $\sigma = 0.2^\circ$). Flanker contrast was identical to that of the Gabor stimuli on each trial. Flankers were tilted either 30° clockwise or counter-clockwise (randomized on each trial) from vertical to perturb responses in the neural subpopulations most sensitive for discriminating the near-vertical target orientation so that performance would benefit by narrowing the attention field. The center-to-center distance between the stimuli and flankers was 3.5°. From the point of view of the model, adding the flankers was the same as increasing stimulus size (beyond the excitatory part of the classical receptive field), but without increasing attention field size (because attending the flankers would have impaired performance). We verified through model

simulations that the magnitude of the response gain change at high contrasts was predicted by the model to be larger with than without the flankers (see also equation (8) of ref. 1), and even larger when both the flankers were present and the attention size was narrowed.

Analysis of psychophysical data. For each observer, performance, $d' = z(\text{hit rate}) - z(\text{false alarm rate})$, was assessed across experimental sessions for each contrast and each trial condition (valid, invalid and neutral). Counter-clockwise response to counter-clockwise stimulus tilt was (arbitrarily) considered to be a hit and counter-clockwise response to clockwise stimulus was considered to be a false alarm. The psychometric data were fit (nonlinear least-squares) to the mean performance (across observers)

$$d'(c) = d'_{\max} \frac{c^n}{c^n + c_{50}^n},$$

where $d'(c)$ is performance as a function of contrast, d'_{\max} determines the asymptotic performance at high contrasts, c_{50} is the contrast corresponding to half the asymptotic performance and n is an exponent that controls the slope of the psychometric function. The two parameters d'_{\max} and c_{50} determined response gain and contrast gain, respectively. We estimated these two parameters for each attention condition while the exponent was treated as one free parameter, constrained to have the same value across conditions (Supplementary Tables 1 and 2). Best-fit parameters were obtained separately for the exogenous and endogenous attention experiments.

A bootstrap procedure was used to determine confidence intervals for the fitted response gain (d'_{\max}) and contrast gain (c_{50}) parameters (Fig. 3). The same bootstrap procedure was used to determine whether changes in response and/or contrast gain were statistically significant. Specifically, we randomly resampled individual psychophysical trials with replacement to generate a resampled dataset, which was refit, and this procedure of resampling and refitting was repeated 10,000 times to generate bootstrap distributions of the psychometric data and of the fitted parameters. Confidence intervals and P values were extracted for each parameter estimate from these bootstrap distributions. We compiled the bootstrap distribution of the differences between the conditions (for example, valid versus invalid trials) and determined the percentage of the values in the tail of the distribution of the differences greater than zero for response gain changes (d'_{\max}), or less than zero for contrast gain changes (c_{50}). The use of these one-tailed statistical tests was justified on the basis of previous studies, reporting a benefit for valid and a cost for invalid, compared to neutral cues^{7–9,29}.

For the individual subject analysis (Fig. 4), we computed the d' performance across experimental sessions separately for each condition and contrast to obtain the psychometric functions. We then followed the same fitting procedures as outlined above to estimate d'_{\max} and c_{50} separately for each observer.

Observers and fMRI scanning sessions. fMRI data were acquired from five healthy observers with normal or corrected vision (25–31 years old, three females), three of whom participated in the psychophysics experiments. Each observer participated in several scanning sessions: one session to obtain a high-resolution anatomical volume for cortical surface extraction and coregistration across sessions, one session to define retinotopically organized visual areas and two to three sessions to measure fMRI responses in the main experiment (spatial uncertainty versus no uncertainty, described below).

MRI data acquisition. MRI data were acquired at 3T (Allegra, Siemens Medical Systems) equipped with a transmit head coil (NM-011) and an eight-channel phased array receive surface coil (NMSC-071; both Nova Medical). We used a standard echoplanar pulse sequence with the following parameters: repetition time = 1.2 s, echo time = 30 ms, flip angle = 72°, voxel size = 3 × 3 × 3 mm, 22 slices oriented approximately parallel to the calcarine sulcus.

Stimulus, task and fMRI procedure. The experimental protocol was similar to the endogenous attention psychophysics experiment (see above). After an initial fixation (250 ms), a pre-cue (250 ms) indicated the upcoming target location. All trials were validly cued (one hemifield was cued, the other uncued) to maximize the number of trials to be included in the analysis. A blank period of 1,250 ms allowed observers to covertly deploy endogenous attention. A pair of Gabor



stimuli (5 cycles per degree, $\sigma = 0.4^\circ$ s.d. of the Gaussian window) were presented (50 ms) in the lower quadrants of the left and right hemifields (6° eccentricity), one of which was the target. The stimuli were tilted slightly clockwise or counter-clockwise of vertical. Stimulus contrast was 10%, the lowest contrast at which observers could perform the task above chance, to minimize the contribution of stimulus-evoked activity to the fMRI measurements. While fixating, observers reported the orientation of the cued stimulus (clockwise or counter-clockwise of vertical) by pressing one of two buttons. Intertrial intervals were variable in duration (5.4–9 s, randomly interleaved).

Two conditions (spatial uncertainty, no spatial uncertainty) were conducted in separate sessions on different days. In one condition, a pair of stimuli (10% contrast) appeared in the left and right hemifields, randomly and independently at one of five predefined locations in each hemifield (see above and Fig. 2c). In the other condition, the stimuli were always shown at the same pair of locations (the middle of the five locations in the first condition) in the left and right hemifields. In this condition, to further minimize spatial uncertainty, placeholders (white line edges of 0.25° length) were displayed continuously throughout each run at the corners of the fixed stimulus locations in each hemifield. Stimulus tilt was adjusted for each observer, such that performance was ~70% correct in both conditions.

fMRI data analysis. Standard retinotopic mapping procedures were used to identify V1, for each observer⁴⁹. Within V1, we defined 14 regions of interest (ROIs, seven in each hemisphere) corresponding to a series of locations along an arc in the visual field, including the five stimulus locations in the lower left and right visual quadrants along with two additional locations in the upper visual field quadrants.

fMRI responses from the attention experiments were preprocessed with standard methods, and a response amplitude was measured for each of the two conditions (spatial uncertainty, no spatial uncertainty), and for each of the ROIs (for details, see ref. 50). In the spatial uncertainty condition, we restricted the analysis to trials in which stimuli were presented at the middle of the five locations. Thus, the responses were measured for stimuli of identical size at the identical location, with and without spatial uncertainty. We then subtracted the response amplitudes

for uncued stimulus presentations from the response amplitudes for cued stimuli presented at the same location. This resulted in a spatial distribution of response differences (that is, the attention field) with one response difference for each of the ROIs, corresponding to a series of locations along an arc in each visual hemifield. Left and right hemisphere responses were analyzed separately and were treated as being statistically independent.

To quantify the attention field size, we computed the spread of the spatial distribution of response differences (cued minus uncued) across the seven ROIs. Specifically, the response differences were treated as a probability distribution and we computed its mean and s.d. The mean was computed as

$$E(x) = \sum x_i p(x_i)$$

where i indexes over the seven ROIs, x_i is a location along the arc in the visual field corresponding to the ROI center and $p(x_i)$ is the cued-uncued response difference for that ROI. Negative values were truncated, and response differences were normalized such that they summed to 1. $E(x)$ therefore represented the center of the attention field and had units of arc degrees of visual angle. The s.d. was computed as

$$Sd(x) = \sqrt{\sum (x_i - E(x))^2 p(x_i)}$$

The size of the attention field was then defined as twice the s.d. in units of arc degree of visual angle in the visual field. This analysis was repeated for both conditions (spatial uncertainty versus no uncertainty), for both hemispheres of each observer (Fig. 5).

47. Liu, T., Stevens, S.T. & Carrasco, M. Comparing the time course and efficacy of spatial and feature-based attention. *Vision Res.* **47**, 108–113 (2007).
48. Sperling, G.D.B.A. (ed.) *Handbook of Perception and Human Performance*, 1–65 (John Wiley & Sons, 1986).
49. Wandell, B.A., Dumoulin, S.O. & Brewer, A.A. Visual field maps in human cortex. *Neuron* **56**, 366–383 (2007).
50. Donner, T.H., Sagi, D., Bonneh, Y.S. & Heeger, D.J. Opposite neural signatures of motion-induced blindness in human dorsal and ventral visual cortex. *J. Neurosci.* **28**, 10298–10310 (2008).